

# Reversible Jump MCMC

Intro to MCMC (Andrieu, et al.)  
Trans-dimensional MCMC (Green)  
Applications of RJMCMC (Laine)

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# Outline

- 1 Overview
  - Problem
  - How to Jump Between Spaces
- 2 Algorithm
  - Detailed balance
  - The Algorithm
- 3 Summary and Discussion

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# Problem Statement

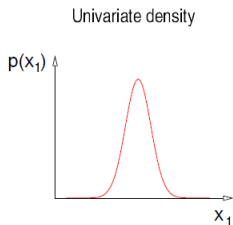
We want to know the probability distribution:

- Exact inference: Sum-product
- Approximation: MCMC
- Approximate for several models: Reverse jump MCMC

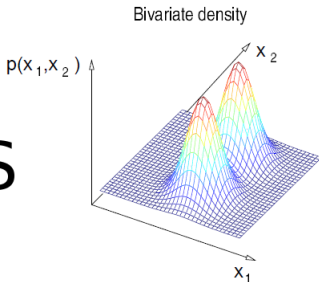
Reverse jump MCMC essentially allows us to jump across different models

“The number of things you don’t know is one of things you don’t know”

Would a better model be to use 1, 2, or 3 parameters?



VS



# Detailed Balance

Recall detailed balance from basic MCMC (discrete)

$$p(x^{(i)})T(x^{(i-1)}|x^{(i)}) = p(x^{(i-1)})T(x^{(i)}|x^{(i-1)})$$

A Markov chain on a continuous general state space

$$\int_{(x,x') \in A \times B} \pi(dx)P(x, dx') = \int_{(x,x') \in A \times B} \pi(dx')P(x', dx).$$

# Detailed Balance

If we want to do Metropolis-Hastings on a general space, we replace the transition kernel by the acceptance probability  $\alpha(x, x')$  and the proposal measure  $q(x, dx')$

$$\int_{(x,x') \in A \times B} \pi(dx) P(x, dx') = \int_{(x,x') \in A \times B} \pi(dx') P(x', dx).$$

$$\int_{(x,x') \in A \times B} \pi(dx) q(x, dx') \alpha(x, x') = \int_{(x,x') \in A \times B} \pi(dx') q(x', dx) \alpha(x', x)$$

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# Defining a General State Space

The space we're sampling from is:

$$X \triangleq \cup_{m=1}^M m \times X_m \quad (2)$$

where

- $x_1 = \{1, 2\}$
- $x_2 = \{(1, 1), (1, 2), (2, 1), (2, 2)\}$
- etc.

$$X = \{(1, 1), (1, 2), (2, (1, 1)), (2, (1, 2)), \dots\} \quad (3)$$

To jump between spaces, we need to create:

- Communicating State Spaces
- Bijective function that maps from one space to the other

# Communicating State Spaces

Dimensions should be equal.

$$\overline{X_{m,n}} \triangleq X_m \times U_{m,n} \quad (4)$$

and

$$\overline{X_{n,m}} \triangleq X_n \times U_{n,m} \quad (5)$$

$U$ 's are random quantities (possibly from a Gaussian distribution).

$X$ 's and  $U$ 's are like vectors (not really)

# Bijjective Function

The function must satisfy these properties:

$$f_{n \rightarrow m}(x_n, u_{n,m}) = (x_m, u_{m,n}) \quad (6)$$

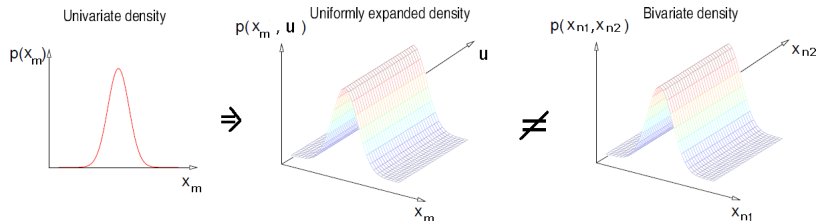
$$f_{m \rightarrow n}(x_m, u_{m,n}) = (x_n, u_{n,m}) \quad (7)$$

which says it must be deterministic and invertible.

$$d(x_m) + d(u_{m,n}) = d(x_n) + d(u_{n,m}) \quad (8)$$

which says the dimensions must be equal.

# Example

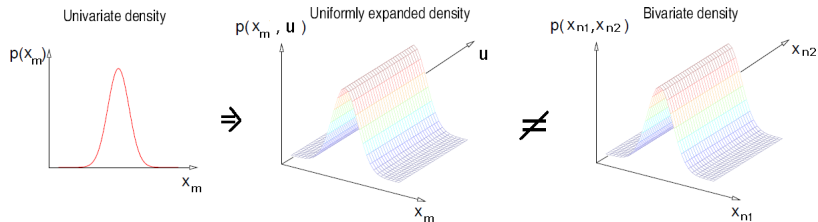


Univariate distribution's area is 1.

Bivariate distribution's volume is 1.

Expanded distribution's volume is 1 (assuming  $u = [0,1]$ )

# Example



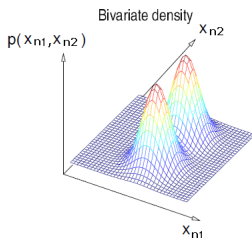
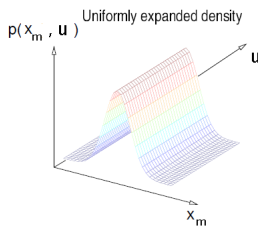
$$(x_m, u) \in X_{m,n}$$

$$(x_{n1}, x_{n2}) \in X_{n,m}$$

$$f_{m \rightarrow n}(x_m, u) = (x_{n1}, x_{n2})$$

$$f_{n \rightarrow m}(x_{n1}, x_{n2}) = (x_m, u)$$

# Example



Most often, we want to capture some information from the other 2 variables into the univariate distribution.

$$f_{m \rightarrow n}(x_m, u) = (x_m + u, x_m - u)$$

$$f_{n \rightarrow m}(x_{n1}, x_{n2}) = \left( \frac{x_{n1} + x_{n2}}{2}, \frac{x_{n1} - x_{n2}}{2} \right)$$

# Example

$$f_{m \rightarrow n}(x_m, u) = (x_m + u, x_m - u)$$

$$f_{n \rightarrow m}(x_{n1}, x_{n2}) = \left( \frac{x_{n1} + x_{n2}}{2}, \frac{x_{n1} - x_{n2}}{2} \right)$$

$$\begin{aligned} f_{m \rightarrow n}(x_m, u) &= (x_m + u, x_m - u) \\ &= \left( \frac{x_{n1} + x_{n2}}{2} + \frac{x_{n1} + x_{n2}}{2}, \frac{x_{n1} + x_{n2}}{2} - \left( \frac{x_{n1} - x_{n2}}{2} \right) \right) \\ &= \left( \frac{2x_{n1}}{2}, \frac{2x_{n2}}{2} \right) \\ &= (x_{n1}, x_{n2}) \end{aligned}$$



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With our previous Metropolis-Hastings on the general state space

$$\int_{(x,x') \in A \times B} \pi(dx) q(x, dx') \alpha(x, x') = \int_{(x,x') \in A \times B} \pi(dx') q(x', dx) \alpha(x', x)$$

$$\alpha(x, x') = \min\left\{1, \frac{\pi(dx') q(x', dx)}{\pi(dx) q(x, dx')}\right\}$$

Substitutes  $q$  and  $\pi$  with  $f$  with respect to  $u$

$$\int_{(x,x') \in A \times B} \alpha(x, x') f(x, x') \mu(dx, dx') = \int_{(x,x') \in A \times B} \alpha(x', x) f(x', x) \mu(dx', dx)$$

$$\alpha(x, x') = \min\left\{1, \frac{f(x', x)}{f(x, x')}\right\} \quad (9)$$

If we were to make random variables  $u$ , our equation looks like

$$\int \pi(x)g(u)\alpha(x, x')dxdu = \int \pi(dx')g(u')\alpha(x', x)dx' du' \quad (10)$$

$g(u)$  : random numbers sampled from known density  $g$

$x' = h(x, u)$  and  $x = h'(x', u')$

We're essentially moving from one state space to another (but with differing number of  $u$ 's).

$$\alpha(x, x') = \min\left\{1, \frac{\pi(x')}{\pi(x)} \times \frac{g'(u')}{g(u)} \times \left| \frac{\partial(x', u')}{\partial(x, u)} \right| \right\}$$

$$\begin{aligned} J_{f_{n \rightarrow m}} &= \left| \det \frac{\partial f_{n \rightarrow m}(x, u)}{\partial (x, u)} \right| \\ &= \det \begin{pmatrix} \frac{\partial x}{\partial x'} & \frac{\partial x}{\partial u'} \\ \frac{\partial u}{\partial x'} & \frac{\partial u}{\partial u'} \end{pmatrix} \end{aligned} \quad (11)$$

Describes a stretching factor or magnification factor from one coordinate system  $(x, y)$  to  $(u, v)$ .

Recall that  $x = h(x', u')$

Modifying the equation to deal with multiple models, we change it to

$$\int \pi(x) q_m(x, x') \alpha_m(x, x') dx du = \int \pi(x') q_m(x', x) \alpha_m(x', x) dx' du'$$

$$\alpha_m = \min\left\{1, \frac{\pi(x')}{\pi(x)} \times \frac{j_m(x')}{j_m(x)} \times \frac{g'_m(u')}{g_m(u)} \times \left| \frac{\partial(x', u')}{\partial(x, u)} \right| \right\}$$

$j_m(x)$  : probability of choosing to move type  $m$  when at  $x$

$g_m(u)$  : random number sampled from known density  $g$

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# Idea

## The Metropolis Hastings algorithm

1. Initialise  $x^{(0)}$ .
2. For  $i = 0$  to  $N - 1$ 
  - Sample  $u \sim \mathcal{U}_{[0,1]}$ .
  - Sample  $x^* \sim q(x^* | x^{(i)})$ .
  - If  $u < \mathcal{A}(x^{(i)}, x^*) = \min \left\{ 1, \frac{p(x^*)q(x^{(i)} | x^*)}{p(x^{(i)})q(x^* | x^{(i)})} \right\}$ 

$$x^{(i+1)} = x^*$$
  - else
 
$$x^{(i+1)} = x^{(i)}$$

Figure 5. Metropolis-Hastings algorithm.

Update  $x$  if the acceptance probability is high enough.

1. Initialisation: set  $(k^{(0)}, \mu^{(0)})$ .
2. For  $i = 0$  to  $N - 1$ 
  - Sample  $u \sim \mathcal{U}_{[0,1]}$ .
  - If  $(u \leq b_k)$ 
    - then “birth” move.
    - else if  $(u \leq b_k + d_k)$  then “death” move.
    - else if  $(u \leq b_k + d_k + s_k)$  then “split” move.
    - else if  $(u \leq b_k + d_k + s_k + m_k)$  then “merge” move.
    - else update.
  - End If.
  - Sample other parameters.

Figure 18. Generic reversible jump MCMC.

Unlike normal MCMC, we can make several decisions depending on our acceptance probability

A generic MJMCMC algorithm has 5 moves:

- Birth: simply add a new component(s)
- Death: simply remove a component(s)
- Split: add a new component by splitting an old one
- Merge: remove 2 components by merging them into one
- Update: Simply update the component

All except update are trans-dimensional

RJMCMC allows us to sample from models of different dimensions.

Issues with RJMCMC include:

- If models were very different, better to do within-model
- Poor performance (?)
- Cumbersome to construct and difficult to tune

# Discussion

Target distribution:

$$p(k, dx) = \sum_{m=1}^M p(m, dx_m) I_{m \times x_m}(k, x) \quad (12)$$