

## Markov chain Monte Carlo methods

- The approximations of expectation so far have assumed that the samples are independent draws.
- This sounds good, but in high dimensions, we do not know how to get **good** independent samples from the distribution.
- MCMC methods drop this requirement.
- Basic intuition
  - If you have **finally** found a region of high probability, stick around for a bit, enjoy yourself, grab some more samples.

## Markov chain Monte Carlo methods

- Samples are conditioned on the previous one (this is the Markov chain).
- MCMC is generally a good hammer for complex, high dimensional, problems.
- Main downside is that it is not “plug-and-play”
  - Doing well requires taking advantage to the structure of your problem
  - MCMC tends to be expensive (but take heart---there may not be any other solution, and at least your problem is being solved).

## Metropolis Example

We want samples  $z^{(1)}, z^{(2)}, \dots$

Again, write  $p(z) = \tilde{p}(z)/Z$

Assume that  $q(z|z^{(prev)})$  can be sampled easily

Also assume that  $q(\cdot)$  is symmetric, i.e.,  $q(z_A|z_B) = q(z_B|z_A)$

For example,  $q(z|z^{(prev)}) \sim \mathcal{N}(z; z^{(prev)}, \sigma^2)$

## Metropolis Example

While not\_bored

{

    Sample  $q(z|z^{(prev)})$

    Accept with probability  $A(z, z^{(prev)}) = \min\left(1, \frac{\tilde{p}(z)}{\tilde{p}(z^{(prev)})}\right)$

    If accept, emit  $z$ , otherwise, emit  $z^{(prev)}$ .

}

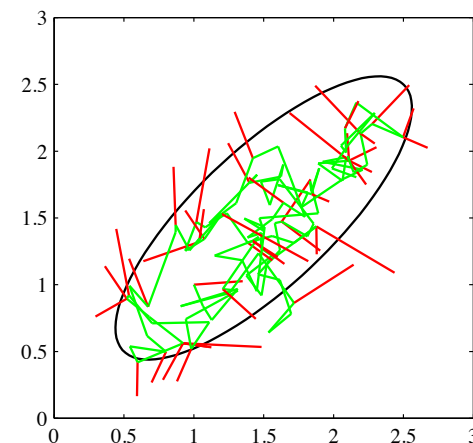
## Metropolis Example

Note that

$$A(z, z^{(prev)}) = \min \left( 1, \frac{\tilde{p}(z)}{\tilde{p}(z^{(prev)})} \right) = \min \left( 1, \frac{p(z)}{p(z^{(prev)})} \right)$$

We do not need to normalize  $p(z)$

## Metropolis Example



Green follows accepted proposals  
Red are rejected moves.

## Markov chain view

Denote an initial probability distribution by  $p(z^{(1)})$

Define transition probabilities by:

$$T(z^{(prev)}, z) = p(z | z^{(prev)}) \quad (\text{a probability distribution})$$

$T = T_m(\ )$  can change over time, but for now, assume that it is always the same (homogeneous chain)

A given chain evolves from a sample of  $p(z^{(1)})$ , and is an instance from an ensemble of chains.

## Stationary Markov chains

- Recall that our goal is to have our Markov chain emit samples from our target distribution.
- This implies that the distribution being sampled at time  $t+1$  is the same as that of time  $t$  (stationary).
- If our stationary (target) distribution is  $p()$ , then if imagine an ensemble of chains, they are in each state with (long-run) probability  $p()$ .
  - On average, a switch from s1 to s2 happens as often as going from s2 to s1, otherwise, the percentage of states would not be stable
- If our stationary (target) distribution is  $p()$ , what do the transition probabilities look like?

## Detailed balance

- Detailed balance is defined by:

$$p(z)T(z, z') = p(z')T(z', z)$$

(We assume that  $T(\bullet) > 0$ )

- Detailed balance is a sufficient condition for a stationary distribution.
- Detailed balance is also referred to as reversibility.

## Detailed balance implies stationary

$$p(z) = \sum_{z'} p(z')T(z', z) \quad (\text{marginalization})$$

If we have detailed balance, then

$$p(z')T(z', z) = p^{(prev)}(z)T(z, z')$$

So,

$$p(z) = \sum_{z'} p(z')T(z', z) = \sum_z p^{(prev)}(z)T(z, z') = p^{(prev)}(z')$$

Hence, detailed balance implies the distribution is stationary.

## Detailed balance (cont)

- Detailed balance (for  $p()$ ) means that *if* our chain was generating samples from  $p()$ , it would continue to do so.
  - We will address how it gets there shortly
- Does the Metropolis algorithm have detailed balance?

## Metropolis Example

While not\_bored

{

Sample  $q(z|z^{(prev)})$

Accept with probability  $A(z, z^{(prev)}) = \min\left(1, \frac{\tilde{p}(z)}{\tilde{p}(z^{(prev)})}\right)$

If accept, emit  $z$ , otherwise, emit  $z^{(prev)}$ .

}

Same as  $\frac{p(z)}{p(z^{(prev)})}$

## Metropolis Example

Recall that in Metropolis,  $A(z, z') = \min\left(1, \frac{p(z)}{p(z')}\right)$

$$\begin{aligned} p(z')q(z|z')A(z, z') &= q(z|z')\min(p(z'), p(z)) \\ &= q(z'|z)\min(p(z'), p(z)) \\ &= p(z)q(z'|z)\min\left(\frac{p(z')}{p(z)}, 1\right) \\ &= p(z)q(z'|z)\min\left(1, \frac{p(z')}{p(z)}\right) \\ &= p(z)q(z'|z)A(z', z) \end{aligned}$$

## Ergodic chains

- Different starting probabilities will give different chains
- We want our chains to converge (in the limit) to the same stationary state, regardless of starting distribution.
- Such chains are called ergodic, and the common stationary state is called the equilibrium state.
- Ergodic chains have a unique equilibrium.

## When do our chains converge?

- Important theorem tells us that (for finite state spaces\*) our chains converge to equilibrium under two relatively weak conditions.
- (1) Irreducible
  - We can get from any state to any other state
- (2) Aperiodic
  - The chain does not get trapped in cycles
- These are true for detailed balance which is sufficient, but not necessary for convergence.

\*Infinite or uncountable state spaces introduces additional complexities.

## Intuition behind ergodic chains

Let  $p^{(t)}(z)$  be the distribution at some time (e.g., initial distribution)

Let  $p^*(z)$  be the stationary distribution

$$\text{Let } p^{(t)}(z) = p^*(z) - q^{(t)}(z)$$

Note that the elements of  $p^{(t+1)}(z)$  and  $p^*(z)$  sum to one, and thus the elements of  $q(z)$  sum to zero.

Note also that  $q(z)$  is not a probability.

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Let  $p^{(t)}(z) = p^*(z) - q^{(t)}(z)$

$$\begin{aligned} p^{(t+1)}(z) &= \sum_{z'} p^{(t)}(z') T(z, z') \\ &= \sum_{z'} p^*(z') T(z, z') - \sum_{z'} q^{(t)}(z') T(z, z') \\ &= p^*(z) - q^{(t+1)}(z) \end{aligned}$$

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Claim that  $|q^{(t+1)}(z)| < |q^{(t)}(z)|$